Universal and automatic elbow detection for learning the effective number of components in model selection problems

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Abstract

We design a universal automatic elbow detector (UAED) for deciding effective number of components in model selection problems. The relationship with the information criteria widely employed in the literature is also discussed. The proposed UAED does not require the knowledge of a likelihood function and can be easily applied in diverse applications, such as regression and classification, feature and/or order selection, clustering, and dimension reduction. Several experiments involving synthetic and real data show the advantages of the proposed scheme with benchmark techniques in the literature.

Keywords: model selection, order selection, automatic elbow detection, variable selection, clustering.

1 Introduction

Model selection is vast and one of the most relevant tasks in signal processing, statistics and machine learning [1, 2, 3]. It is the process of selecting a statistical model from a set of candidate ones. Model selection includes as special cases very famous sub-tasks: order selection (e.g., in polynomial functions or ARMA models [4]), variable selection [5], dimension reduction [6], and clustering [7], to name a few.

More specifically, in a large amount of research works from the most diverse fields, researchers and practitioners face a trade-off between the number of components/variables to consider in their analyses and the performance of the obtained results. Note that we use the term "variables" as a general concept that can equivalently represent variables, features, or number of clusters, depending on the nature of the considered problem. This trade-off occurs because increasing the number of variables taken into account in the analysis allows for better results, at the expense of obtaining a more complex model. In other words, the model performance and the model complexity generate the so-called bias-variance trade-off. Therefore, in many applications, researchers must obtain the optimal number of components/variables to take into account the aforementioned trade-off [1].

The solution in the literature belongs to different families and approaches. A first class of methods

is formed by the resampling techniques, such as cross-validation (CV) or bootstrap, where the dataset is split into training and test sets [8, 9, 10]. However, the proportion of data to include in the training and test sets is a crucial parameter that affects critically the results. Another important family is the class of the *information criteria* [11], such as the Bayesian information criterion (BIC) [12], the Akaike information criterion (AIC) [13], or the Hannan-Quinn information criterion (HQIC) [14], to name a few [2, 15]. The information criteria consider a linear penalization of the model complexity, and they differ for the choice of the slope of this penalization. These choices are motivated by theoretical probabilistic derivations which involve several assumptions and approximations. Hence, the good performance of an information criterion is often restricted to very specific scenarios. Moreover, the computation of the information criteria often requires the knowledge of the maximum of a likelihood function. Other probabilistic strategies related to the information criteria are the so-called minimum description length principle, Mallows's Cp coefficient and the structural risk minimization [16, 17]. In the Bayesian framework, the use of marginal likelihood and posterior predictive approaches are usually employed [2, 18, 19]. The connection between the marginal likelihood and information criteria is discussed in the appendices of [15]. The posterior predictive approach is related to the CV idea. Furthermore, standard frequentist approaches based on p-values have a vast use in some specific applications and deserve to be cited [20, 21]. Finally, specially in the clustering literature, some authors apply a visual inspection of an error curve looking for an "elbow".

In this work, we design an universal automatic elbow detector (UAED) based on a geometric approach. The proposed scheme is inspired by the concept of the maximum "area under the curve" (AUC) in receiver operator characteristic (ROC) curves [1, 22], which is well-known and vastly employed in signal processing and machine learning. The resulting UAED technique also induces a linear penalization of the model complexity. We discuss the connections, differences and the advantages of UAED with respect to the information criteria already presented in the literature. It is important to remark that the range of applicability of UAED is much wider than other techniques in the literature, since no likelihood function is required. The application of UAED only requires the knowledge of an error curve, that can be defined in different ways according to the user's needs. Moreover, we describe several appealing behaviours of the UAED and test it in different numerical examples, two of them involving a real dataset. The results show the benefits of UAED with respect to other benchmark techniques in the literature.

The remainder of the article is organised as follows: Section 2 describes the framework and the notation employed in the development of the UAED, Section 3 presents and discusses the UAED in detail, and Section 4 shows some UAED tests in practical applications. Finally, in Section 5 conclusions are given.

2 Framework and main notation

In many applications, we desire to infer a vector of parameters $\boldsymbol{\theta}_k = [\theta_1, ..., \theta_k]^{\top}$ of dimension k given a data vector $\mathbf{y} = [y_1, ..., y_N]^{\top}$. A likelihood function $p(\mathbf{y}|\boldsymbol{\theta}_k)$ is usually available, often induced by a related physical model. Furthermore, in different types of real world application problems (clustering, variable selection, or dimension reduction) and specially in model selection

problems, an error function (i.e., a fitting measure) is obtained, that we denote as

$$V(k): \mathbb{N} \to \mathbb{R}, \qquad k = 0, 1, 2, ..., K,$$

where k denotes the number of components (e.g., variables, clusters, or order of the polynomial function), i.e., k defines the complexity of the model. In the literature, we often have

$$V(k) = -2\log(\ell_{\max}), \quad \text{ where } \quad \ell_{\max} = \max_{\boldsymbol{\theta}} p(\mathbf{y}|\boldsymbol{\theta}_k),$$

as in [11]. However, in this work, V(k) could-be directly the mean square error (MSE), or the mean absolute error (MAE). For instance, V(k) can represent the prediction error in regression problems with a polynomial function, where k is the order of the polynomial, or the sum of the inner variances within clusters where k is the number of clusters. We assume that k starts in 0 and grows with step 1 for simplicity, but more general cases can easily be addressed.

Generally, V(k) is a non-increasing error curve, i.e., for any pair of non-negative integers n_1, n_2 such that $n_2 > n_1$, then we have $V(n_2) \leq V(n_1)$. Indeed, V(k) is a fitting term that decreases as the complexity of the model (given by the number k of parameters) grows. Therefore, we have

$$V(0) \ge V(k), \quad \forall k.$$

Observe that V(0) represents the value of the error function corresponding, for instance, to a constant model in a regression problem, or a single cluster (for all the data) in a clustering problem. See Figure 1(a) for a graphical example of the curve V(k). In some applications, the score function V(k) should be also convex, i.e., the differences V(n+1) - V(n) will decrease as n increases. This is the case of a variable selection problem, if the variables have been ranked correctly. However, this work does not require conditions regarding the concavity of V(k).

Additional assumptions. Just for the sake of simplicity and without loss of generality, we assume that $\min V(k) = V(K) = 0$. Note that this condition can be always obtained with a simple subtraction, defining a new curve $V'(k) = V(k) - \min V(k) = V(k) - V(K)$. Moreover, above we have assumed k = 0, 1, ..., K but, if there exists a value $k_{\text{max}} \leq K$ such that V(k) has not an additional drop for $k \geq k_{\text{max}}$, i.e.,

$$V(k_{\max}) = V(k_{\max} + 1) = V(k_{\max} + 2) = \dots = V(K), \tag{1}$$

in this scenario, we can consider $k=0,1,...,k_{max}$, since the rest of the components must be discarded because they do not cause a drop in the error function. See Figures 1(a)-1(b) for two graphical examples. Clearly, if the minimum value of k is different from 0, let us say k_{min} , we can always set $k'=k-k_{min}$. Finally, a different value of the incremental step can be also considered. So far, we have assumed an incremental step of 1, without loss of generality.

¹This condition could be also relaxed. We keep it, for the sake of simplicity.

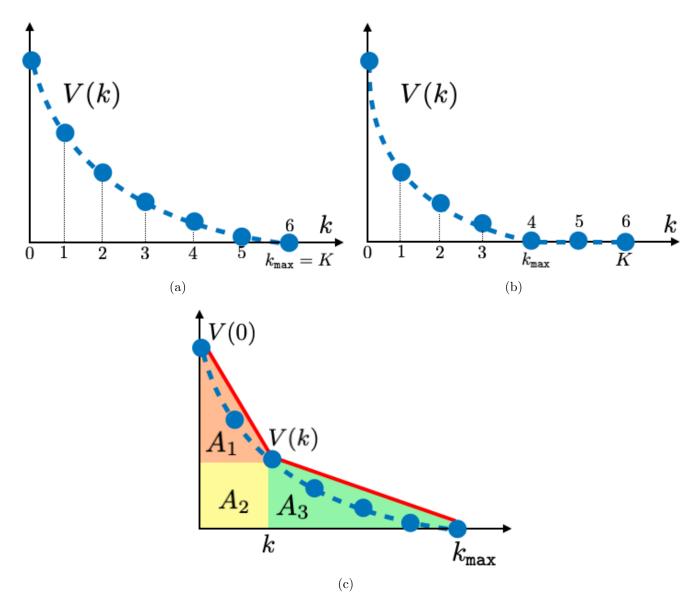


Figure 1: (a)-(b) Example of error curve V(k) where (a) $k_{\text{max}} = K = 6$, (b) $k_{\text{max}} = 4$ and K = 6. (c) Construction with two straight lines and the areas A_1 , A_2 and A_3 .

3 The Universal Automatic Elbow Detector (UAED)

In this section, we provide two equivalent geometric derivations of the proposed method, and discuss the similarities, differences, and connections with other methods in the literature. The behaviour of the proposed technique is described and some interesting considerations are also highlighted.

3.1 First derivation

Considering the decay V(k) described in the previous section, the underlying idea is "inspired" by the concept of the maximum AUC in ROC curves [1, 22]. Namely, we desire to extract geometric information from the curve V(k) looking for an "elbow" in order to determine the optimal number of components, denoted $k^* \in \{0, 1..., k_{max}\}$, to consider in our model (i.e., in the vector $\boldsymbol{\theta}_{k^*}$). We consider the construction of two straight lines passing through the points (0, V(0)), (k, V(k)) and (k, V(k)), $(k_{max}, 0)$ as shown in Figure 1(c) (where $k \in \{0, 1, ..., k_{max}\}$). These two straight lines form a piece-wise linear approximation of the curve V(k). The goal is to minimize the area under this approximation. More specifically, as we can see in Figure 1(c), the area to minimize consist of three sub-areas: two areas of two triangles $(A_1 \text{ and } A_3)$ and the area of a rectangle in the middle (A_2) . Namely, we have

$$A_{1} = \frac{k \cdot (V(0) - V(k))}{2},$$

$$A_{2} = k \cdot V(k),$$

$$A_{3} = \frac{(k_{\text{max}} - k) \cdot V(k)}{2},$$
(2)

hence the definition of k^* is

$$k^* = \arg\min_{k} \{ A_1 + A_2 + A_3 \},$$

$$= \arg\min_{k} \left\{ \frac{V(k)}{V(0)} + \frac{k}{k_{\text{max}}} \right\}, \quad \text{for } k = 1, ..., k_{\text{max}}.$$
(3)

Multiplying by the constant value V(0), we can equivalently write

$$k^* = \arg\min_{k} \left\{ V(k) + \frac{V(0)}{k_{\text{max}}} k \right\}, \quad \text{for } k = 1, ..., k_{\text{max}}.$$
 (4)

It is important to remark that, since k belongs to a discrete and finite set, solving the optimization above is straightforward (if K, or k_{max} , is not a huge value). In the case of multiple minima, e.g., having M different minima, $k_1^*, k_2^*, ..., k_M^*$, the user can choose the best solution (within the M possible one) according to some specific requirement depending on the specific application. Here, we suggest the most conservative choice, i.e.,

$$k^* = \max k_j^*. (5)$$

3.2 Second equivalent derivation

The solution offered by the expressions (3)-(4) is equivalent to finding the k^* such that the difference between $V(k^*)$ and the value of the straight line (evaluated at k^* , as well) connecting the extreme points (0, V(0)) and $(k_{\text{max}}, 0)$ is maximized, as depicted in Figure 2. More specifically, this straight line has an equation

$$v(k) = -\frac{V(0)}{k_{\text{max}}} \cdot k + V(0),$$

hence, the difference that we maximize is the following:

$$d(k) = v(k) - V(k), \tag{6}$$

$$= -\frac{V(0)}{k_{\text{max}}} \cdot k + V(0) - V(k), \tag{7}$$

$$=V(0)-\left(\frac{V(0)}{k_{\max}}\cdot k+V(k)\right). \tag{8}$$

Since V(0) does not depend on k (i.e., it is a constant value), we can write

$$k^* = \arg\max_k d(k) = \arg\max_k \left[V(0) - \left(\frac{V(0)}{k_{\text{max}}} \cdot k + V(k) \right) \right], \tag{9}$$

$$= \arg\max_{k} \left[-\left(\frac{V(0)}{k_{\max}} \cdot k + V(k) \right) \right], \tag{10}$$

$$= \arg\min_{k} \left[\frac{V(0)}{k_{\text{max}}} \cdot k + V(k) \right], \tag{11}$$

which is exactly the expression in Eq. (4). Two additional and equivalent derivations are given in Appendix A and Appendix B. They are also represented graphically in Figures 5(a) and 5(b), respectively.

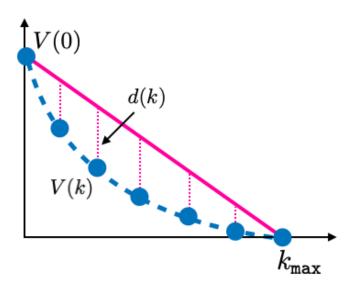


Figure 2: Graphical representation of alternative derivation in Section 3.2.

3.3 Relation with the information criteria

Recalling the expression in (4), i.e.,

$$k^* = \arg\min\left\{V(k) + \frac{V(0)}{k_{\max}}k\right\}.$$

here we show that this cost function can be interpreted in the same form of other information criteria, i.e., with a linear penalization of the model complexity,

$$C(k) = V(k) + \frac{V(0)}{k_{\text{max}}}k, \tag{12}$$

$$=V(k)+\lambda k,\tag{13}$$

where we set $\lambda = \frac{V(0)}{k_{\text{max}}}$. Note that Eq. (13) has exactly the same form of the cost function used in the information criteria like BIC and AIC, for instance, when V(k) is defined as

$$V(k) = -2 \log \ell_{\max}, \quad \text{ with } \quad \ell_{\max} = \max_{\boldsymbol{\theta}} p(\mathbf{y}|\boldsymbol{\theta}_k).$$

BIC corresponds to the choice $\lambda = \log(N)$ where N is the number of data in \mathbf{y} , and AIC corresponds to the choice $\lambda = 2$. Therefore, when $V(k) = -2\log\ell_{\max}$, UAED can be interpreted as an information criterion with the particular choice of $\lambda = \frac{V(0)}{k_{\max}}$. Table 1 summarizes this information.

Table 1: Different information criteria and the proposed UAED.

Criterion	Choice of λ
Bayesian-Schwarz information criterion (BIC) [12]	$\log N$
Akaike information criterion (AIC) [13]	2
Hannan-Quinn information criterion (HQIC) [14]	$\log(\log(N))$
Universal Automatic Elbow Detector (UAED)	$rac{V(0)}{k_{ exttt{max}}}$

3.4 Behaviour of the proposed solution

Analyzing the involved parameters in the expression (4) or (12), we can highlight the following considerations about the behaviour of the UAED method. We list some important points below:

• Observing Eq. (12), the penalization of the complexity of the model depends on V(0) and k_{max} : since $\lambda = \frac{V(0)}{k_{\text{max}}}$ increasing V(0) or decreasing k_{max} , intensifies the penalty. This is a reasonable and desirable behaviour. Indeed, increasing the value of V(0) also increases the differences V(0)-V(k), which means that the first components/variables have more impact in the fitting - the decay of V(k) - so that fewer components/variables can form a reasonable model. Otherwise, decreasing the value of V(0) means more variables have a similar impact in the decay of V(k). Therefore,

we should consider more components, in fact the slope of the penalization, $\lambda = \frac{V(0)}{k_{\text{max}}}$, decreases in this case.

- Regarding k_{max} , we can notice that a decrease of k_{max} means that fewer components/variables produces a drop in the curve V(k). On the other hand, an increase in k_{max} means that the use of more variables causes a drop V(k), so we should consider more components, indeed, the slope of the penalization, $\lambda = \frac{V(0)}{k_{\text{max}}}$, decreases.
- Looking the expression (4) or (12), it is possible to show that the solution does not depend on different possible re-normalization of the axes, i.e., scaling of the axes the solution remains invariant (one of them, or both, even with different scales). Indeed, considering a scaling on the vertical axis, i.e., assuming V(k)' = aV(k) with a > 0, we have

$$\begin{split} k^* &= \arg\min\left[aV(k) + \frac{aV(0)}{k_{\max}}k\right] = \arg\min\left[a\left(V(k) + \frac{V(0)}{k_{\max}}k\right)\right], \\ &= \arg\min\left[V(k) + \frac{V(0)}{k_{\max}}k\right]. \end{split}$$

Let now consider the case of scaling the horizontal axis, for instance, instead of having $k = 0, 1, 2..., k_{\text{max}}$, we have $k' = 0, b, 2b, ..., bk_{\text{max}}$ (i.e., k' = bk), and another error curve $\widetilde{V}(k') = V(k'/b)$, where b is a positive integer. Hence we can write

$$\begin{split} (k')^* &= \arg\min_{k'} \left[\widetilde{V}(k') + \frac{V(0)}{k'_{\max}} \cdot k' \right] = \arg\min_{k'} \left[V(k'/b) + \frac{V(0)}{k'_{\max}} \cdot k' \right] = \\ &= b \arg\min_{k} \left[V(bk/b) + \frac{V(0)}{bk_{\max}} \cdot bk \right] \\ &= b \arg\min_{k} \left[V(k) + \frac{V(0)}{k_{\max}} \cdot k \right] = bk^*. \end{split}$$

Namely, the new solution $(k')^* = bk^*$ is just a scaled version of the previous one, taking into account the factor b. Furthermore, given the considerations in Appendices A and B, we can see that the solution is invariant even if the axes are exchanged.

• Here, we describe two ideal scenarios and discuss the behaviour of UAED. For clarity in the exposition, let us consider as an example a variable selection problem. First of all, we consider the case that all the input variables are equally important for predicting the output variable. Then, we have $k_{\text{max}} = K$, and the error curve V(k) is a straight line connecting the points (0, V(0)) and $(k_{\text{max}}, V(k_{\text{max}}))$ (i.e., each variable has the same impact to the error decay). In this scenario, we have $k_{\text{max}} = K$, and UAED provides $M = k_{\text{max}} + 1$ different minima $k_1^* = 0, k_2^* = 1, k_3^* = 2, ..., k_M^* = k_{\text{max}}$. Thus, the UAED solution is given by Eq. (5), i.e., $k^* = \max k_j^* = k_{\text{max}}$. Namely, UAED suggests to select all the variables, that is the correct solution.

On the other hand, let us consider now a scenario where all the input variables are independent from the output variable. In this case, V(k) is a constant function, i.e., V(k) = V(0) for all k and, as a consequence, $k_{\text{max}} = 0$. Hence, since $k_{\text{max}} = 0$, UAED gives $k^* = 0$, which is the correct solution (i.e., no variables should be selected). Thus, in both scenarios, UAED provides the correct results.

4 Experiments with synthetic and real data

In this section, we test the UAED in four real-world applications. In each experiment, we consider a different function V(k), in order to show the vast range of applicability of UAED. Sections 4.1-4.2 deal with synthetic data in clustering and order selection problems, respectively. In Sections 4.3-4.4, the experiments involve the analysis of real data: the first one is a variable selection in a regression problem with soundscape emotion data and, the second one is a classification problem with biomedical data. We compare the performance of UAED with BIC, AIC, and other information criteria described in the literature, in the examples where these schemes can be also applied.

4.1 Clustering

We consider 2500 simulated data from a mixture of 5 bidimensional Gaussian distributions, $\mathcal{N}(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)$, where $\boldsymbol{\mu}_1 = [3,0], \ \boldsymbol{\Sigma}_1 = [0.3,0;0,2], \ \boldsymbol{\mu}_2 = [14,5], \ \boldsymbol{\Sigma}_2 = [1.5,0.7;0.7,1.5];, \ \boldsymbol{\mu}_3 = [-5,-10], \ \boldsymbol{\Sigma}_3 = [1.5,0.7;0.7,1.5], \ \boldsymbol{\mu}_4 = [10,-10], \ \boldsymbol{\Sigma}_4 = [1.5,0;0,1.5];, \ \text{and} \ \boldsymbol{\mu}_5 = [-5,5], \ \boldsymbol{\Sigma}_5 = [1,-0.8;-0.8,1].$ Figure 3(a) shows these data points.

We assume $V(k) = \log \left[\sum_{j=1}^{k+1} \operatorname{var}(j) \right]$, where $\operatorname{var}(j)$ represents the inner variance of the j-th cluster, as shown in Figure 3(b). Each value of $\operatorname{var}(j)$ has been computed and averaged over 200 runs, applying a k-means algorithm. In this setting, the total number of clusters is given by k+1 (i.e., k=0 corresponds to a unique, single cluster). We assume K=50 as the maximum number of possible clusters.

It is important to remark that, with this choice of V(k), the other information criteria cannot be directly applied². We apply UAED and obtain $k^* + 1 = 5$ as the chosen number of clusters, which is the correct solution.

4.2 Order selection of a polynomial function in a regression problem

We generate a dataset of N = 100 pairs $\{x_n, y_n\}_{n=1}^N$, where both inputs x_n 's and outputs y_n 's are scalar values, considering the following observation model,

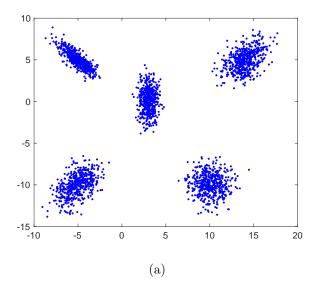
$$y_n = \theta_0 + \theta_1 x_n + \theta_2 x_n^2 + \dots \theta_k x_n^k + \epsilon_n, \tag{14}$$

where $\boldsymbol{\theta}_k = [\theta_0, \theta_1, ..., \theta_k]^{\top}$, ϵ_n is Gaussian noise with zero mean and variance $\sigma_{\epsilon}^2 = 1$. The dataset has been generated with a polynomial function of order k = 4, and with the coefficients

$$\theta_0 = 4.05, \ \theta_1 = -2.025, \ \theta_2 = -2.225, \ \theta_3 = 0.1, \ \theta_4 = 0.1.$$

In this experiment, we consider $V(k) = -2\log(\ell_{\text{max}})$ with $\ell_{\text{max}} = \max_{\boldsymbol{\theta}} p(\mathbf{y}|\boldsymbol{\theta}_k)$ with $k \leq K$, where $p(\mathbf{y}|\boldsymbol{\theta}_k)$ is induced by Eq. (14), in order to allow the comparison with other schemes in the literature, as shown in Table 1. The corresponding function V(k) is shown in Figure 4(a).

The information criteria require the choice of the error curve of type $V(k) = -2 \log \ell_{\text{max}}$ where $\ell_{\text{max}} = \max_{\boldsymbol{\theta}} p(\mathbf{y}|\boldsymbol{\theta}_k)$ and, as a consequence, a definition of a likelihood function $p(\mathbf{y}|\boldsymbol{\theta})$.



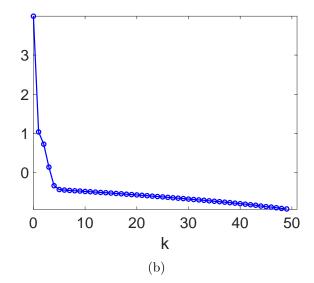


Figure 3: (a) Artificial Data of the clustering experiment. (b) The function $V(k) = \log \left[\sum_{j=1}^{k+1} \text{var}(j) \right]$ where var(j) represents the inner variance in the j-th cluster. Note that k = 0 corresponds to a unique, single cluster.

Applying BIC, AIC and Hannan-Quinn IC we obtain the suggested order of polynomial is 4, 6, and 10, respectively. With the proposed UAED method, we obtain the suggested order is 4, which is the correct order of the underlying polynomial function. Therefore, in this experiment, BIC and UAED provide the correct answer.

4.3 Variable selection in a regression problem with real data

In this section, we present a feature selection problem for regression. Moreover, we consider real data. More specifically, a dataset of N pairs $\{\mathbf{x}_n, y_n\}_{n=1}^N$ is given, where each input vector $\mathbf{x}_n = [x_{n,1}, ..., x_{n,K}]$ is formed by K variables, and the outputs y_n 's are scalar values. We assume $K \leq N$ and a linear observation model,

$$y_n = \theta_0 + \theta_1 x_{n,1} + \theta_2 x_{n,2} + \dots \theta_K x_{n,K} + \epsilon_n, \tag{15}$$

where ϵ_n is Gaussian noise with zero mean and variance σ_{ϵ}^2 , i.e., $\epsilon_n \sim \mathcal{N}(\epsilon|0, \sigma_{\epsilon}^2)$. In the real dataset studied in [23], there are K = 122 features and N = 1214 number of data points. The output represents the variable defined as "arousal" in [23].

In order to allow the comparison with other schemes in the literature, here we can set $V(k) = -2\log(\ell_{\text{max}})$ where $\ell_{\text{max}} = \max_{\theta} p(\mathbf{y}|\boldsymbol{\theta}_k)$ with $k \leq K$, after ranking the 122 variables as in [23]. Clearly, The likelihood function $p(\mathbf{y}|\boldsymbol{\theta}_k)$ is induced by Eq. (15). Therefore, in this experiment, we can compare UAED again with other information criterion measures in the literature, some of them are given in Table 1. BIC suggests a model with 17 variables, AIC chooses 44 variables, the Hannan-Quinn IC selects 41 variables. The proposed UAED suggests considering only 11 variables, Therefore, the UAED suggestions is closer to the results given in other previous studies and to experts' recommendations in the literature, e.g., [23].

4.4 Variable selection in a classification problem with real data

The authors in [24] analyze the most important features for predicting patients at risk of developing nonalcoholic fatty liver disease. The authors collected data from 1525 patients who attended the Cardiovascular Risk Unit of Mostoles University Hospital (Madrid, Spain) from 2005 to 2021, and use a random forest (RF) algorithm to classify patients and rank the input features, in order to select the most important one. They found that 4 features were the most relevant according to the ranking and the experts' opinions: (a) insulin resistance, (b) ferritin, (c) serum levels of insulin, and (d) triglycerides.

In this experiment, we set $V(k) = 1 - \operatorname{accuracy}(k)$ that is given in Figure 4(b), after ranking the 35 features [24]. Note that V(0) = 0.5 representing a completely random binary classification. It is important to remark that, with this choice of V(k), the other information criteria cannot be employed.² The application of UAED suggests to select 4 variables which is exactly the result of the paper [24], obtained using a Cross-Validation (CV) approach, and supported by the experts' opinions.

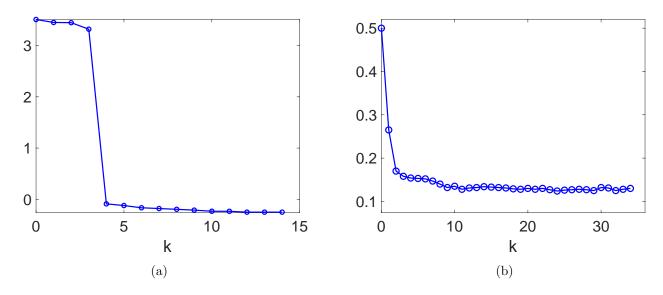


Figure 4: (a) The corresponding curve $V(k) = -2 \log \ell_{\text{max}}$ (with $\ell_{\text{max}} = \max_{\theta} p(\mathbf{y}|\boldsymbol{\theta}_k)$) in Section 4.2; (b) The curve V(k) = 1 - accuracy(k) of the experiment in Section 4.4.

5 Conclusions

A novel universal automatic elbow detector (UAED) has been introduced. Four different geometrical derivations have been provided. Moreover, we have analyzed its behavior and properties, as the invariance on scaling the axes and the behavior in ideal scenarios. The relationships and differences with other information criteria (already given in the literature) have been described and highlighted. Furthermore, the proposed procedure has a much wider

range of application with respect to the other schemes in the literature. Several experiments and comparisons show the benefits of the proposed UAED scheme.

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References

- [1] C. M. Bishop, "Pattern recognition," Machine Learning, vol. 128, pp. 1–58, 2006.
- [2] F. Llorente, L. Martino, D. Delgado, and J. Lopez-Santiago, "Marginal likelihood computation for model selection and hypothesis testing: an extensive review," (to appear) SIAM Review arXiv:2005.08334, 2022.
- [3] J. Ding, V. Tarokh, and Y. Yang, "Model selection techniques: An overview," *IEEE Signal Processing Magazine*, vol. 35, no. 6, pp. 16–34, 2018.
- [4] P. Stoica and Y. Selen, "Model-order selection: a review of information criterion rules," *IEEE Signal Processing Magazine*, vol. 21, no. 4, pp. 36–47, 2004.
- [5] V. Bolón-Canedo, N. Sánchez-Maroño, and A. Alonso-Betanzos, "A review of feature selection methods on synthetic data," *Knowledge and information systems*, vol. 34, no. 3, pp. 483–519, 2013.
- [6] Y. Ma and L. Zhu, "A review on dimension reduction," *International Statistical Review*, vol. 81, no. 1, pp. 134–150, 2013.
- [7] A. Saxena, M. Prasad, A. Gupta, N. Bharill, O. P. Patel, A. Tiwari, M. J. Er, W. Ding, and C.-T. Lin, "A review of clustering techniques and developments," *Neurocomputing*, vol. 267, pp. 664–681, 2017.
- [8] P. Stoica and Y. Selén, "Cross-validation rules for order estimation," *Digital Signal Processing*, vol. 14, pp. 355–371, 2004.
- [9] E. Fong and C. Holmes, "On the marginal likelihood and cross-validation," *Biometrika*, vol. 107, no. 2, pp. 489–496, 2020.
- [10] A. Vehtari, A. Gelman, and J. Gabry, "Practical Bayesian model evaluation using leave-one-out cross-validation and WAIC," Statistics and computing, vol. 27, no. 5, pp. 1413–1432, 2017.
- [11] S. Konishi and G. Kitagawa, *Information criteria and statistical modeling*. Springer Science & Business Media, 2008.

- [12] G. Schwarz *et al.*, "Estimating the dimension of a model," *The annals of statistics*, vol. 6, no. 2, pp. 461–464, 1978.
- [13] D. Spiegelhalter, N. G. Best, B. P. Carlin, and A. V. der Linde, "Bayesian measures of model complexity and fit," *J. R. Stat. Soc. B*, vol. 64, pp. 583–616, 2002.
- [14] E. J. Hannan and B. G. Quinn, "The determination of the order of an autoregression," *Journal of the Royal Statistical Society. Series B (Methodological)*, vol. 41, no. 2, pp. 190–195, 1979.
- [15] F. Llorente, L. Martino, E. Curbelo, J. Lopez-Santiago, and D. Delgado, "On the safe use of prior densities for bayesian model selection," WIREs Computational Statistics, p. e1595, 2022.
- [16] M. Kobayashi and S. Sakata, "Mallows' cp criterion and unbiasedness of model selection," *Journal of Econometrics*, vol. 45, no. 3, pp. 385–395, 1990.
- [17] J. Shawe-Taylor, P. L. Bartlett, R. C. Williamson, and M. Anthony, "Structural risk minimization over data-dependent hierarchies," *IEEE transactions on Information Theory*, vol. 44, no. 5, pp. 1926–1940, 1998.
- [18] J. A. Hoeting, D. Madigan, A. E. Raftery, and C. T. Volinsky, "Bayesian model averaging: a tutorial," *Statistical Science*, vol. 14, no. 4, pp. 382–417, 1999.
- [19] C. M. Pooley and G. Marion, "Bayesian model evidence as a practical alternative to deviance information criterion," *Royal Society Open Science*, vol. 5, no. 3, pp. 1–16, 2018.
- [20] M. Efroymson, "Multiple regression analysis," Mathematical methods for digital computers, pp. 191–203, 1960.
- [21] R. R. Hocking, "The analysis and selection of variables in linear regression," *Biometrics*, pp. 1–49, 1976.
- [22] J. A. Hanley and B. J. McNeil, "The meaning and use of the area under a receiver operating characteristic (ROC) curve," *Radiology*, vol. 143, no. 1, pp. 29–36, 1982.
- [23] R. San Millán-Castillo, L. Martino, E. Morgado, and F. Llorente, "An exhaustive variable selection study for linear models of soundscape emotions: Rankings and Gibbs analysis," *IEEE/ACM Transactions on Audio, Speech, and Language Processing*, vol. 30, pp. 2460–2474, 2022.
- [24] R. Gárcia-Carretero, R. Holgado-Cuadrado, and O. Barquero-Pérez, "Assessment of classification models and relevant features on nonalcoholic steatohepatitis using random forest," *Entropy*, vol. 23, no. 6, 2021.

Third alternative derivation A

Let us consider Figure 5(a). First of all, we must find the value k' such that the straight line, connecting the points (0, V(0)) and $(k_{\text{max}}, 0)$, reaches the value V(k) (where $k \neq k'$, and more precisely $k \leq k'$). Namely, we desire to obtain k' such that

$$V(k) = -\frac{V(0)}{k_{\text{max}}} \cdot k' + V(0),$$

hence

$$k' = -\frac{k_{\text{max}}}{V(0)} \left[V(k) - V(0) \right].$$

Now, we could also consider to maximize the following difference

$$r(k) = k' - k, (16)$$

$$= -\frac{k_{\text{max}}}{V(0)} \left[V(k) - V(0) \right] - k, \tag{17}$$

and the elbow is defined as

$$k^* = \arg\max \ r(k) = \arg\max \left[-\frac{k_{\text{max}}}{V(0)}V(k) - k \right], \tag{18}$$

$$= \arg\min\left[\frac{k_{\max}}{V(0)}V(k) + k\right],\tag{19}$$

$$= \arg\min\left[V(k) + \frac{V(0)}{k_{\text{max}}}k\right],\tag{20}$$

where in the last we have multiplied by the constant V(0). Note that Eq. (20) is exactly the same optimization problem (i.e., with the same cost function) in Sections 3.1-3.2.

\mathbf{B} Fourth alternative derivation

One could also consider the Euclidean distance e(k) between the points in the curve V(k) and the straight line connecting the points (0, V(0)) and $(k_{\text{max}}, 0)$, as depicted in Figure 5(b). Observing this figure, we can notice that

$$e(k) = d(k)\sin(\pi/2 - \alpha) = d(k)\cos\alpha \tag{21}$$

$$e(k) = r(k)\sin\alpha,\tag{22}$$

where α is the angle shown in Figure 5(b). Since the angle α is constant, then we can write

$$k^* = \arg\max_k e(k) = \arg\max_k d(k) \cos \alpha = \arg\max_k d(k)$$

$$= \arg\max_k r(k) \sin \alpha = \arg\max_k r(k).$$
(23)

$$= \arg\max_{k} r(k) \sin \alpha = \arg\max_{k} r(k). \tag{24}$$

Therefore, maximizing e(k) is equivalent to maximize d(k) or r(k).

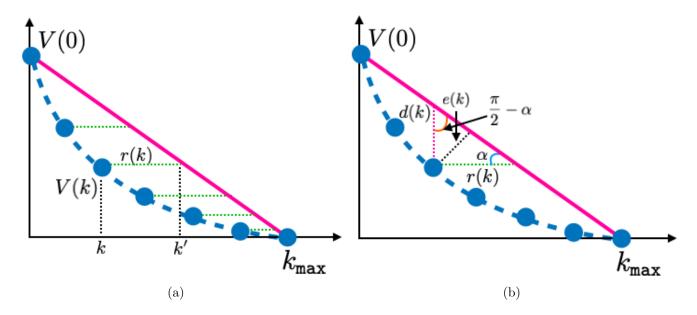


Figure 5: (a) Graphical representation of the other alternative derivation in Appendix A. (b) Graphical representation of derivation based on the Euclidean distance e(k).

C Possible extension

We have already shown that the resulting expression in Eq. (3) provides good performance and is endowed with valuable behaviours.

However, we can add more flexibility that can be useful in the scenarios in which the researchers and/or practitioners determine that the benefit of reducing the error is greater than the benefit of reducing the number of considered variables or vice versa. We define an additional parameter $\alpha \in [0, 1]$, and consider the modified definition of the optimal k as

$$k^* = \arg\min_{k} \left[\alpha \cdot \frac{V(k)}{V(0)} + (1 - \alpha) \cdot \frac{k}{k_{\text{max}}} \right]. \tag{25}$$

Note that $\alpha = 0$ implies that all priority is to reduce the number of considered variables $(k^* = 0)$, that $\alpha = 1$ implies that all priority is to reduce the resulting error (so that $k^* = k_{\text{max}}$). For $\alpha = 0.5$, we come back to the definition in Eq. (4). As we have previously done in Section 3, can rewrite Eq. (25) as

$$k^* = \arg\min_{k} \left[V(k) + \underbrace{\left(\frac{1 - \alpha V(0)}{\alpha k_{\text{max}}}\right) \cdot k}_{\lambda} \right], \qquad (26)$$

$$k^* = \arg\min_{k} \left[V(k) + \lambda k \right],$$

having the form of an information criterion with a different choice of λ which involves now the parameter α , as well.