

Estimation of Mean in Presence of Non Response Using Exponential Estimator

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Abstract

This paper considers the problem of estimating the population mean using information on auxiliary variable in presence of non response. Exponential ratio and exponential product type estimators have been suggested and their properties are studied. An empirical study is carried out to support the theoretical results.

Keywords: population mean, study variable, auxiliary variable, exponential ratio, exponential, product estimator, Bias, MSE

1. Introduction

In surveys covering human populations, information is in most cases not obtained from all the units in the survey even after some call-backs. Hansen and Hurwitz(1946) considered the problem of non response while estimating the population mean by taking a sub sample from the non respondent group with the help of some extra efforts and an estimator was proposed by combining the information available from response and non response groups. In estimating population parameters like the mean, total or ratio, sample survey experts sometimes use auxiliary information to improve precision of the estimates. When the population mean \bar{X} of the auxiliary variable x is known and in presence of non response, the problem of estimation of population mean \bar{Y} of the study variable y has been discussed by

Cochran (1977), Rao (1986), Khare and Srivastava(1997) and Singh and Kumar (2008). In Hansen and Hurwitz(1946) method, questionnaires are mailed to all the respondents included in a sample and a list of non respondents is prepared after the deadline is over. Then a sub sample is drawn from the set of non respondents and a direct interview is conducted with the selected respondents and the necessary information is collected.

Assume that the population is divided into two groups, those who will not respond called non-response class. Let N_1 and N_2 be the number of units in the population that belong to the response class and the non-response class respectively ($N_1+N_2=N$). Let n_1 be the number of units responding in a simple random sample of size n drawn from the population and n_2 the number of units not responding in the sample. We may regard the sample of n_1 respondents as a simple random sample from the response class and the sample of n_2 as a simple random sample from the non-response class. Let h_2 denote the size of the sub sample from n_2 non-respondents to be interviewed and $f = \frac{n_2}{h_2}$. Let \bar{y}_1 and \bar{y}_{h_2} denote the sample means of y character based on n_1 and h_2 units respectively. The estimator proposed by Hansen and Hurwitz (1946) is given by-

$$\bar{y}^* = \frac{n_1 \bar{y}_1 + n_2 \bar{y}_{h_2}}{n} \quad (1.1)$$

The estimator \bar{y}^* is unbiased and has variance

$$V(\bar{y}^*) = \left(\frac{1}{n} - \frac{1}{N}\right) S_y^2 + (f-1) \frac{N_2}{N} \cdot \frac{S_{y_2}^2}{n} \quad (1.2)$$

The population mean square of the character y is denoted by S_y^2 and the population mean square of y for N_2 non-response units of the population is denoted by $S_{y_2}^2$.

Bahl and Tuteja (1991) introduced an exponential ratio –type estimator for population mean as given by

$$\bar{y}_{er} = \bar{y} \exp \left[\frac{\bar{X} - \bar{x}}{\bar{X} + \bar{x}} \right] \quad (1.3)$$

and exponential product- type estimator as

$$\bar{y}_{ep} = \bar{y} \exp \left[\frac{\bar{X} - \bar{x}}{\bar{X} + \bar{x}} \right] \quad (1.4)$$

The objective of this paper is to study Bahl and Tuteja (1991) exponential ratio- type and product- type estimators in presence of non-response.

2. Suggested estimator

First we assume that the non response is only on study variable. The estimator \bar{y}_{er} under non response will take the form

$$\bar{y}_{er}^* = \bar{y}^* \exp \left[\frac{\bar{X} - \bar{x}}{\bar{X} + \bar{x}} \right] \quad (2.1)$$

To obtain the bias and MSE of the estimator \bar{y}_{er}^* we write

$$\bar{y}^* = \bar{Y}(1 + e_0^*), \quad \bar{x} = \bar{X}(1 + e_1)$$

Such that

$$E(e_0^*) = E(e_1) = 0,$$

and

$$\left. \begin{aligned} E(e_0^2) &= \frac{V(\bar{y}^*)}{\bar{Y}^2} \\ E(e_1^2) &= \frac{V(\bar{x})}{\bar{X}^2} = \left(\frac{1}{n} - \frac{1}{N}\right) \frac{S_x^2}{\bar{X}^2} \\ E(e_0^* e_1) &= \frac{\text{cov}(\bar{y}^*, \bar{x})}{\bar{Y}\bar{X}} = \left(\frac{1}{n} - \frac{1}{N}\right) \frac{S_{xy}}{\bar{X}\bar{Y}} \end{aligned} \right\} \quad (2.2)$$

where $S_{xy} = \frac{1}{(N-1)} \sum_{j=1}^N (x_j - \bar{X})(y_j - \bar{Y})$.

Now expressing \bar{y}_{er}^* in terms of e's we have

$$\bar{y}_{er}^* = \bar{Y}(1 + e_0) \exp \left\{ \frac{-e_1}{2\left(1 + \frac{e_1}{2}\right)} \right\} \quad (2.3)$$

Expanding the right hand side of (2.3) and neglecting the terms involving powers of e's greater than two we have

$$\bar{y}_{er}^* = \bar{Y} \left(1 + e_0 - \frac{e_1}{2} - \frac{e_0 e_1}{2} + \frac{3e_1^2}{8} \right) \quad (2.4)$$

Taking expectations of both sides of (2.4), we get the bias of \bar{y}_{er}^* to the first degree of approximation, as

$$B(\bar{y}_{er}^*) = \left(\frac{1}{n} - \frac{1}{N}\right) \bar{Y} \left[\frac{3C_x^2}{8} - \frac{\rho C_y C_x}{2} \right] \quad (2.5)$$

Squaring both sides of (2.4) and neglecting terms of e's involving powers greater than two we have

$$(\bar{y}_{er}^* - \bar{Y})^2 = \bar{Y}^2 \{e_0^2 + \frac{e_1^2}{4} - e_0 e_1\} \quad (2.6)$$

Taking expectations of both sides of (2.6) we get the MSE (to the first degree of approximation) as

$$MSE(\bar{y}_{er}^*) = (\frac{1}{n} - \frac{1}{N}) \bar{Y}^2 [C_y^2 + \frac{C_x^2}{4} - \rho C_y C_x] + \frac{(f-1)N_2}{nN} S_{y2}^2 \quad (2.7)$$

3. Exponential product type estimator

The estimator \bar{y}_{ep} under non response (only on study variable) will take the form

$$\bar{y}_{ep}^* = \bar{y}^* \exp\left[\frac{\bar{x} - \bar{X}}{\bar{x} + \bar{X}}\right] \quad (3.1)$$

Following the procedure of section 2, we get the bias and MSE of \bar{y}_{ep}^* as

$$Bias(\bar{y}_{ep}^*) = (\frac{1}{n} - \frac{1}{N}) \bar{Y} \left[\frac{-C_x^2}{8} + \frac{\rho C_x C_y}{2} \right] \quad (3.2)$$

and

$$MSE(\bar{y}_{ep}^*) = (\frac{1}{n} - \frac{1}{N}) \bar{Y}^2 [C_y^2 + \frac{C_x^2}{4} + \rho C_y C_x] + \frac{(f-1)N_2}{nN} S_{y2}^2 \quad (3.3)$$

4. Non response on both y and x

We assume that the non response is both on study and auxiliary variable. The estimator \bar{y}_{er} and \bar{y}_{ep} under non response on both the variables takes the following form respectively-

$$\bar{y}_{er}^{**} = \bar{y}^* \exp\left[\frac{\bar{X} - \bar{x}^*}{\bar{X} + \bar{x}^*}\right] \quad (4.1)$$

$$\bar{y}_{ep}^{**} = \bar{y}^* \exp \left[\frac{\bar{x}^* + \bar{X}}{\bar{x}^* + \bar{X}} \right] \quad (4.2)$$

To obtain the bias and MSE of the estimator \bar{y}_{er}^{**} and \bar{y}_{ep}^{**} we write

$$\bar{y}^* = \bar{Y}(1 + e_0), \quad \bar{x}^* = \bar{X}(1 + e_1^*), \quad \bar{x} = \bar{X}(1 + e_1)$$

Such that

$$E(e_0) = E(e_1^*) = 0$$

$$\left. \begin{aligned} E(e_1^{*2}) &= \frac{v(\bar{x}^*)}{\bar{X}^2} = \left(\frac{1}{n} - \frac{1}{N} \right) S_x^2 + (f-1) \frac{N_2}{N} \cdot \frac{S_{x2}^2}{n} \\ E(e_0 e_2) &= \frac{\text{cov}(\bar{y}^*, \bar{x}^*)}{\bar{Y}\bar{X}} \\ &= \frac{1}{\bar{Y}\bar{X}} \left[\left(\frac{1}{n} - \frac{1}{N} \right) S_{xy} + (f-1) \frac{N_2}{N} \cdot \frac{S_{xy}}{n} \right] \end{aligned} \right\} \quad (4.3)$$

The population mean square of the character x is denoted by S_x^2 and the population mean square of x for N_2 non response units of the population is denoted by S_{x2}^2 .

The biases and MSE of the estimators \bar{y}_{er}^{**} and \bar{y}_{ep}^{**} are given by \bar{y}_{er}^{**} and \bar{y}_{ep}^{**}

$$\begin{aligned} B(\bar{y}_{er}^{**}) &= \left(\frac{1}{n} - \frac{1}{N} \right) \bar{Y} \left[\frac{3}{8} \frac{S_x^2}{\bar{X}^2} - \frac{S_{xy}}{2\bar{Y}\bar{X}} \right] + \bar{Y} \frac{(f-1)}{n} \cdot \frac{N_2}{N} \left[\frac{3}{8} \frac{S_{x2}^2}{\bar{X}^2} - \frac{S_{xy2}}{2\bar{Y}\bar{X}} \right] \\ &= \left(\frac{1}{n} - \frac{1}{N} \right) \bar{Y} \left[\frac{3}{8} C_x^2 - \frac{\rho C_x C_y}{2} \right] + \bar{Y} \frac{(f-1)}{n} \cdot \frac{N_2}{N} \left[\frac{3}{8} C_x'^2 - \frac{\rho_2 C_x' C_y'}{2} \right] \end{aligned} \quad (4.4)$$

where $C'_x = \frac{S_{x2}}{\bar{X}}$, $C'_y = \frac{S_{y2}}{\bar{Y}}$.

$$\begin{aligned} B(\bar{y}_{ep}^{**}) &= \left(\frac{1}{n} - \frac{1}{N}\right)\bar{Y}\left[\frac{-1}{8}\cdot\frac{S_x^2}{\bar{X}^2} + \frac{S_{xy}}{2\bar{Y}\bar{X}}\right] + \bar{Y}\frac{(f-1)}{n}\cdot\frac{N_2}{N}\left[\frac{-1}{8}\cdot\frac{S_{x2}^2}{\bar{X}^2} + \frac{S_{xy2}}{2\bar{Y}\bar{X}}\right] \\ &= \left(\frac{1}{n} - \frac{1}{N}\right)\bar{Y}\left[\frac{-1}{8}\cdot C_x^2 + \frac{\rho C_x C_y}{2}\right] + \bar{Y}\frac{(f-1)}{n}\cdot\frac{N_2}{N}\left[\frac{-1}{8}\cdot C_x'^2 + \frac{\rho_2 C_x' C_y'}{2}\right] \end{aligned} \quad (4.5)$$

$$\begin{aligned} \text{MSE}(\bar{y}_{er}^{**}) &= \left(\frac{1}{n} - \frac{1}{N}\right)\bar{Y}^2\left[C_y^2 + \frac{C_x^2}{4} - \rho C_x C_y\right] \\ &+ \bar{Y}^2\frac{(f-1)}{n}\cdot\frac{N_2}{N}\left[C_y'^2 + \frac{C_x'^2}{4} - \rho_2 C_x' C_y'\right] \end{aligned} \quad (4.6)$$

$$\begin{aligned} \text{MSE}(\bar{y}_{ep}^{**}) &= \left(\frac{1}{n} - \frac{1}{N}\right)\bar{Y}^2\left[C_y^2 + \frac{C_x^2}{4} + \rho C_x C_y\right] \\ &+ \bar{Y}^2\frac{(f-1)}{n}\cdot\frac{N_2}{N}\left[C_y'^2 + \frac{C_x'^2}{4} + \rho_2 C_x' C_y'\right] \end{aligned} \quad (4.7)$$

From expressions (2.7),(3.3),(4.6),(4.7), we observe that the MSE expressions of suggested estimators have an additional term (which depends on non-response)as compared to the estimator proposed by Bahl and Tuteja ((1991)(without non response).

5. Efficiency comparisons:

From (1.2), (2.7), (3.3), (4.6) and (4.7), we have

First we compare the efficiencies of \bar{y}_{er}^* and \bar{y}^*

$$\text{MSE}(\bar{y}_{er}^*) - V(\bar{y}^*) \leq 0$$

$$\frac{C_x^2}{4} - \rho C_x C_y \leq 0$$

$$\rho \geq \frac{C_x}{4C_y} \quad (5.1)$$

When this condition is satisfied \bar{y}_{er}^* will be better estimator than \bar{y}^* .

Next we compare the efficiencies of \bar{y}_{ep}^* and \bar{y}^*

$$MSE(\bar{y}_{ep}^*) - V(\bar{y}^*) \leq 0$$

$$\frac{C_x^2}{4} + \rho C_x C_y \leq 0$$

$$\rho \leq -\frac{C_x}{4C_y} \quad (5.2)$$

When this condition is satisfied \bar{y}_{ep}^* will be better estimator than \bar{y}^* .

Next, we compare the efficiencies of \bar{y}_{er}^{**} and \bar{y}^*

$$MSE(\bar{y}_{er}^{**}) - V(\bar{y}^*) \leq 0$$

$$(N - n)\alpha \leq (1 - f)N_2\alpha'$$

$$\frac{\alpha}{\alpha'} \leq \frac{(1 - f)N_2}{N - n} \quad (5.3)$$

$$\text{where } \alpha = \frac{C_x^2}{4} - \rho C_y C_x, \quad \alpha' = \frac{C_x^2}{4} - \rho C_y' C_x'$$

When this condition is satisfied \bar{y}_{er}^{**} will be better estimator than \bar{y}^* .

Finally we compare the efficiencies of \bar{y}_{ep}^{**} and \bar{y}^*

$$MSE(\bar{y}_{ep}^{**}) - V(\bar{y}^*) \leq 0$$

$$(N - n)\lambda \leq (1 - f)N_2\lambda'$$

$$\frac{\lambda}{\lambda'} \leq \frac{(1 - f)N_2}{N - n} \quad (5.4)$$

$$\text{where } \lambda = \frac{C_x^2}{4} + \rho C_y C_x, \quad \lambda' = \frac{C_x^2}{4} + \rho C'_y C'_x$$

When this condition is satisfied \bar{y}_{ep}^{**} will be better estimator than \bar{y}^* .

6. Empirical study

For numerical illustration we consider the data used by Khare and Sinha (2004,p.53). The values of the parameters related to the study variate y (the weight in kg) and the auxiliary variate x (the chest circumference in cm) have been given below.

$$\bar{Y} = 19.50 \quad \bar{X} = 55.86 \quad S_y = 3.04$$

$$S_x = 3.2735 \quad S_{y2} = 2.3552 \quad S_{x2} = 2.51$$

$$\rho = 0.85 \quad \rho_2 = 0.7290$$

$$N_1 = 71 \quad N_2 = 24 \quad N = 95 \quad n = 35$$

Here, we have computed the percent relative efficiencies (PRE) of different suggested estimators with respect to usual unbiased estimator \bar{y}^* for different values of f.

Table 5.1: PRE of different proposed estimators

Values of w	f values	\bar{y}^*	\bar{y}_{er}^*	\bar{y}_{ep}^*	\bar{y}_{er}^{**}	\bar{y}_{ep}^{**}
0.10	1.50	100	263.64	45.47	263.65	45.47
	2.00	100	263.62	45.47	263.65	45.47
	2.50	100	263.61	45.48	263.65	45.47
	3.00	100	263.59	45.48	263.64	45.47
0.20	1.50	100	263.62	45.47	263.65	45.47
	2.00	100	263.59	45.48	263.64	45.47
	2.50	100	263.56	45.48	263.64	45.47
	3.00	100	263.52	45.48	263.63	45.47
0.30	1.50	100	263.61	45.48	263.65	45.47
	2.00	100	263.56	45.48	263.64	45.47
	2.50	100	263.51	45.48	263.63	45.47
	3.00	100	263.46	45.48	263.62	45.47

From table 5.1, we conclude that the estimators which use auxiliary information performs better than Hansen and Hurwitz(1946) estimators \bar{y}^* . Also when non response rate increases, the efficiencies of suggested estimators decreases.

References:

- Bahl, S. and Tuteja, R.K. (1991): Ratio and product type exponential estimator. *Infrm. and Optim.Sci.*,XII,I,159-163.
- Cochran, W.G. (1977): *Sampling techniques*, 3rd edn. Wiley, New York.
- Hansen, M.H. and Hurwitz, W.N. (1946): The problem of non-response in sample surveys. *J. Am. Stat. Assoc.* 41:517–529.
- Khare B.B. and Srivastava, S. (1997): Transformed ratio type estimators for the population mean in the presence of non-response. *Commun. Stat. Theory Methods* 26(7):1779–1791.
- Khare, B.B. and Sinha, R.R. (2004): Estimation of finite population ratio using two-phase sampling scheme in the presence of non-response. *Aligarh J Stat*, 24:43–56.
- Rao, P.S.R.S.(1986): Ratio estimation with sub-sampling the non-respondents. *Surv Methodol* 12(2):217–230.
- Singh, H.P. and Kumar, S. (2008): Estimation of mean in presence of non-response using two phase sampling scheme. *Statistical Papers*.